Delayed paternal age of reproduction in humans is associated with longer telomeres across two generations of descendants

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Telomeres are repeating DNA sequences at the ends of chromosomes that protect and buffer genes from nucleotide loss as cells divide. Telomere length (TL) shortens with age in most proliferating tissues, limiting cell division and thereby contributing to senescence. However, TL increases with age in sperm, and, correspondingly, offspring of older fathers inherit longer telomeres. Using data and samples from a longitudinal study from the Philippines, we first replicate the finding that paternal age at birth is associated with longer TL in offspring (\(n = 2,023, P = 1.84 \times 10^{-6}\)). We then show that this association of paternal age with offspring TL is cumulative across multiple generations: in this sample, grandchildren of older paternal grandfathers at the birth of fathers have longer telomeres (\(n = 234, P = 0.038\), independent of, and additive to, the association of their father's age at birth with TL. The lengthening of telomeres predicted by each year that the father's or grandfather's reproduction are delayed is equal to the yearly shortening of TL seen in middle-age to elderly women in this sample, pointing to potentially important impacts on health and the pace of senescent decline in tissues and systems that are cell-replication dependent. This finding suggests a mechanism by which humans could extend late-life function as average age at reproduction is delayed within a lineage.

adaptation | epigenetics | evolution | parental effects | transgenerational plasticity

Telomeres are repeating DNA sequences at the ends of chromosomes that protect and buffer genes from nucleotide loss as cells divide (1). In many tissues, telomere lengths (TL) are shortened by cellular proliferation, and as a result TL tends to decline with age (2–5). As cell replication generally requires a minimal TL, shortened TL is thought to contribute to senescence (6). Consistent with this, elderly persons with shorter telomeres (in blood) for their age have reduced survival (7–13).

Although it is well established that TL shortens with age in most proliferating tissues (e.g., 4, 5), sperm TL is an exception—older men have sperm with longer telomeres (4, 14, 15). This may be explained by the fact that the activity of telomerase (an enzyme that extends TL) is high in testes (16, 17). Consistent with this, elderly men have sperm with longer telomeres (4, 18, 19). In contrast, because the pool of ova is established in utero, TL in ova are thought to be stable with age, and there is no evidence for a maternal age effect on TL in offspring (e.g., 4, 20).

We recently hypothesized that the age-related TL increase in sperm could lead to cumulative, and thus more biologically significant, multigenerational lengthening or shortening of TL in response to population trends in reproductive scheduling (21). If average reproductive age of recent patrilineal ancestors cumulatively influences TL, this might also lead to changes in TL of sufficient magnitude to influence late-life function and life expectancies (17). To test the hypothesis that a man’s age at reproduction influences TL in grandchildren, we used a large longitudinal, multigeneration sample from Cebu, Philippines (22) in which we measured TL (23) of DNA extracted from venous blood. We related TL in mothers (aged 36–69 y at blood collection) to their fathers’ ages when the mothers were born. In their offspring (21–23 y at the time of blood collection), we related TL to their fathers’ ages when the offspring were born and to their grandfathers’ ages when their parents were born.

Results

Consistent with the expectations of age-related decline in TL, blood TL was inversely associated with age in the 36- to 69-y-old mothers (Fig. L4) with a similar magnitude of effect as seen in previous comparable studies (4, 24). Longer age-adjusted maternal TL predicted longer offspring TL (Fig. 1B). This mother–offspring TL correlation was similar in magnitude to previous studies (25–28) and did not vary depending on the sex of the offspring (maternal age adjusted TL × offspring sex interaction, \(P = 0.81\)). The TL of male offspring (\(n = 902\), mean TL = 0.763 ± 0.006) was on average shorter than the TL of female offspring (\(n = 820\), mean TL = 0.792 ± 0.006, \(t = 3.52, df = 1,720, P = 0.0005\) (Fig. 1B).

As observed in previous studies, we found that longer TL was predicted by older paternal age at reproduction in both the offspring cohort (Table 1, model 1, \(n = 1,681, P = 4.00 \times 10^{-5}\)) and their mothers (Table 2, model 1, \(n = 342, P = 0.003\); mother and offspring combined: \(n = 2,023, P = 1.84 \times 10^{-6}\)). The association of paternal age with offspring TL was little changed by controlling for birth order, household income, or body mass index (BMI) at the time of blood collection in both the offspring (Table 1, model 2) and their mothers (Table 2, model 2). The paternal age association did not differ in sons versus daughters (paternal age × sex interaction, \(P = 0.844\)). Consistent with a linear association, paternal age did not exhibit a quadratic relationship with offspring TL (paternal age\(^2\), \(P = 0.689\) for offspring and \(P = 0.995\) for mothers).

The association of paternal age with longer TL is thought to be due to direct inheritance from longer TL in sperm (4). To test an alternative hypothesis that offspring of older fathers have a slower age-related attrition rate of TL in adulthood, we examined if the association between offspring age and TL was reduced among offspring of older fathers, but found no support for this hypothesis (paternal age × offspring age interaction term in offspring, \(P = 0.341\) and \(P = 0.472\) in mothers).

Author contributions: D.T.A.E., M.G.H., and C.W.K. designed research; D.T.A.E. performed research; D.T.A.E. analyzed data; and D.T.A.E., M.G.H., and C.W.K. wrote the paper.

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\(^{2}\)M.G.H. and C.W.K. contributed equally to this work.
Consistent with our hypothesis, we found that, independent of the association of TL with the father’s age at reproduction, the paternal grandfather’s age at reproduction predicted longer TL in grandchildren (Table 1, model 4, \( n = 234, P = 0.038 \)). The association between TL and grandpaternal age was nearly identical in size to the association with paternal age (Fig. 2; Table 1, models 1 and 4; Table 3), and was little changed after adjusting for household income and BMI at the time of blood collection (Table 1, model 5). As was found for the association of paternal age with TL, the relationship between paternal grandfather’s age and TL did not vary by the age of the grandchild (grand-paternal age \( \times \) grandchild age interaction, \( P = 0.920 \)) or by the sex of grandchild (grand-paternal age \( \times \) sex interaction, \( P = 0.370 \)) and did not exhibit a quadratic relationship (grand-paternal age \( \times \) grandchild age interaction, \( P = 0.809 \)).

The age of the maternal grandfather at the birth of the mother was also positively associated with TL, but this relationship was not significant (Table 1, model 7). In accord with the fact that male gametes continue to divide throughout life, whereas female gametes are produced prenatally, TL did not vary in relation to maternal age at the time of birth: after adjusting for paternal and grandpaternal age, maternal and grandmaternal ages were not significantly associated with the child’s or the grandchild’s TL (Table 1, models 3, 6, and 8; Table 2, model 3).

To place the magnitude of the paternal age associations with TL in a biological context, we compared the increase in TL predicted by each year that the father or grandfather delayed reproduction.

![Fig. 1](image)

**Table 1.** Linear regression models relating paternal age and grandparental age to telomere lengths (relative T/S ratios) in 21- to 23-y-old offspring (both sexes pooled)

<table>
<thead>
<tr>
<th>Model 1</th>
<th>Model 2</th>
<th>Model 3</th>
<th>Model 4</th>
<th>Model 5</th>
<th>Model 6</th>
<th>Model 7</th>
<th>Model 8</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age (self)(^1)</td>
<td>-0.048**</td>
<td>-0.056**</td>
<td>-0.048**</td>
<td>0.0068</td>
<td>-0.010</td>
<td>-0.0013</td>
<td>0.022</td>
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<tr>
<td>Male</td>
<td>-0.031***</td>
<td>-0.035***</td>
<td>-0.031***</td>
<td>-0.055*</td>
<td>-0.061**</td>
<td>-0.054*</td>
<td>-0.063***</td>
</tr>
<tr>
<td>Male ( \times ) age (self)(^1)</td>
<td>-0.030</td>
<td>-0.025</td>
<td>-0.030</td>
<td>-0.071</td>
<td>-0.062</td>
<td>-0.086</td>
<td>-0.13*</td>
</tr>
<tr>
<td>Father’s age</td>
<td>0.0027***</td>
<td>0.0030**</td>
<td>0.0029**</td>
<td>0.0035(^1)</td>
<td>0.0034</td>
<td>0.0041</td>
<td></td>
</tr>
<tr>
<td>Birth order</td>
<td>-0.0019</td>
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<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Family income (log)</td>
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<td></td>
<td></td>
<td></td>
<td></td>
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<tr>
<td>BMI</td>
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<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mother’s age</td>
<td>-0.00030</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Father’s father’s age</td>
<td>0.0029*</td>
<td>0.0026(^a)</td>
<td>0.0029</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Father’s mother’s age</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mother’s father’s age</td>
<td>0.0011</td>
<td>-0.00061</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mother’s mother’s age</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Constant</td>
<td>0.72***</td>
<td>0.67***</td>
<td>0.72***</td>
<td>0.62***</td>
<td>0.52***</td>
<td>0.61***</td>
<td>0.76***</td>
</tr>
<tr>
<td>Observations</td>
<td>1,681</td>
<td>1,602</td>
<td>1,681</td>
<td>234</td>
<td>218</td>
<td>194</td>
<td>342</td>
</tr>
<tr>
<td>Adjusted ( R^2 )</td>
<td>0.052</td>
<td>0.034</td>
<td>0.031</td>
<td>0.045</td>
<td>0.048</td>
<td>0.038</td>
<td>0.047</td>
</tr>
</tbody>
</table>

\(^{1}\)Age is mean centered. \(^{a}\)Values are \( \beta \)-coefficients; \( ^{P} < 0.10 \); \( ^{*} P < 0.05 \); \( ^{**} P < 0.01 \); and \( ^{***} P < 0.001 \).
to the yearly shortening of TL in the mothers (age range: 36–69 y). The yearly decline in TL in the mothers was −0.0032 y (Table 2, model 1; Table 3), which is equal in magnitude, but opposite in direction, to the increase in TL predicted by a 1-y delay in reproduction by the father or paternal grandfather.

**Discussion**

In this large sample from the Philippines, we found that telomeres measured in blood were longer in individuals whose paternal grandfathers were older at their father’s birth. This effect was equal in size, and additive to, the longer TL predicted by a delay in the father’s age at their own birth, suggesting that paternal age at reproduction has effects on TL that are transmitted with high fidelity and cumulatively across at least two generations. The longer TL predicted by each year that paternal or grandpaternal reproduction was delayed was comparable to the yearly shortening in TL seen in older individuals in this population, suggesting that these intergenerational changes may be biologically important. These findings point to a mechanism by which shifts in the age of paternal reproduction within a population can lead to cumulative, multigenerational changes in telomere length in descendants.

Consistent with previous studies (4, 18, 19), the increase in TL predicted by each year that the father’s reproduction was delayed was nearly identical in size to the yearly age-related decline in TL in middle- to old-age adults. The lack of interaction between paternal age and offspring age in models predicting offspring TL, as well as the similar effect sizes of paternal age on TL in the younger and older cohorts (offspring and their mothers, respectively), suggests that the effect of paternal age on TL does not act by changing the rate of telomere shortening in adulthood. However, because both the cohorts were already in adulthood, we are not able to rule out the possibility that the pace of telomere shortening in somatic tissues varies in relation to paternal or grandpaternal age at conception for those ages younger than that of the mothers and offspring in this study. Some (4, 19), but not all (18) previous studies have suggested that the father’s age at conception has greater effects on TL in male offspring. In Cebu, we found that the change in TL predicted by a delay in the father’s or grandfather’s ages at birth did not vary by offspring sex.

A standard model of genetic inheritance leads to the prediction that each offspring receives half of his or her telomere DNA from their father and one quarter of their telomere DNA from each of their paternal and maternal grandfathers. On the basis of this, we expected the effect of delaying reproduction among paternal and maternal grandfathers to be equal in size and half as strong as the effect of a comparable delay in paternal age at reproduction. Contrary to this expectation, we found that the effect size of the paternal grandfather’s age on their grandchildren’s TL was larger than the effect of the maternal grandfather’s age at the mother’s birth on the grandchild’s TL. In addition, the paternal age association with offspring’s TL and the paternal grandfather age association with the grandchild’s TL were equal. Although it is not certain what accounts for these results, both findings might be explained by the fact that heritability of TL is substantially greater from father to offspring than from mother to offspring (25–27), which recent work suggests could relate to telomere protection proteins that are specific to sperm (29, 30).

The longer TL among offspring of older fathers has been explained by the fact that sperm TL are longer in older men (4, 14, 15), with high telomerase activity in testes being the most common explanation for the increase in sperm TL seen with age (16, 31). Because men produce an estimated 100 million sperm cells daily (32–34), telomere maintenance mechanisms are required to avoid rapid telomere shortening (17). It is presently less clear why high testicular telomerase expression should lead to gradual and progressive lengthening of sperm telomeres with age, rather than simply maintaining a stable length. If testicular telomerase expression extends short and long telomeres equally, the distribution of sperm TL in young and old men would be expected to have the same shape, but centered on a longer mean length in older men. However, Kimura and colleagues noted that, compared with younger men, older men have a TL distribution skewed slightly in favor of long telomeres (4). Telomerase preferentially extending long telomeres might explain this, but to the extent that telomerase has effects that vary based upon the length of telomeres, past work suggests that it is likely to have the strongest effects on short telomeres (reviewed in ref. 21). Kimura and colleagues interpreted the skewed distribution of sperm TLs in older men as evidence that sperm stem cells with shorter telomere lengths disproportionately die out with age (4). It is presently not clear if this selective cellular attrition would be sufficient to account for the age-related rate of telomere lengthening observed in sperm.

To our knowledge, no study has used longitudinal sampling of sperm or studied full siblings to firmly establish whether the longer TL in offspring of older male patrilineal ancestors is

**Table 2. Linear regression models relating paternal age to telomere lengths (relative T/S ratios) in 36- to 69-y-old mothers**

<table>
<thead>
<tr>
<th>Model</th>
<th>Model 1</th>
<th>Model 2</th>
<th>Model 3</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age (self)</td>
<td>−0.0032*</td>
<td>−0.0031*</td>
<td>−0.0018</td>
</tr>
<tr>
<td>Father’s age</td>
<td>0.0032**</td>
<td>0.0033*</td>
<td>0.0024</td>
</tr>
<tr>
<td>Family income (log)</td>
<td>−0.0082</td>
<td>0.0016</td>
<td></td>
</tr>
<tr>
<td>BMI</td>
<td>0.0016</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mother’s age</td>
<td>0.0026</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Constant</td>
<td>0.70***</td>
<td>0.70***</td>
<td>0.60***</td>
</tr>
<tr>
<td>Observations</td>
<td>342</td>
<td>339</td>
<td>290</td>
</tr>
<tr>
<td>Adjusted $R^2$</td>
<td>0.030</td>
<td>0.030</td>
<td>0.041</td>
</tr>
</tbody>
</table>

Values are β-coefficients; *P < 0.10 (none); **P < 0.05; ***P < 0.01; and 0.01 ≤ P < 0.001.
caused by increases in sperm TL as a man ages. Although longer sperm TL in older men compared with younger men (4, 14, 15) provides convergent evidence for this prominent hypothesis, it is also known that sperm donors represent nonrandom samples of the population (35, 36). It is possible that older men who volunteer to donate sperm have better reproductive health and longer sperm TL and that a similar selection bias results in men with longer sperm TL tending to reproduce at later ages. Such a strong and persistent selection bias would imply important effects of TL on male reproductive health or that some unmeasured factor influences both sperm TL and reproductive health (which could make sperm TL an important biomarker).

Regardless of the mechanism underlying the longer telomeres in descendants of older fathers and grandfathers, there is evidence that TL is a predictor of health and mortality, suggesting that intergenerational lengthening or shortening of TL could have effects on health and the pace of senescence (17). Several human studies report higher mortality among elderly individuals with shorter blood telomere lengths (7–13). There are converging lines of evidence that telomere lengths per se are at least partial causes of health and survival. Because the mechanics of DNA replication lead to incomplete replication of chromosome ends (the “end-replication problem”), the buffer of repeating telomere sequence can be depleted after multiple rounds of cell division. If replication continues past this point, genes may be progressively lost. To prevent gene loss, shortened telomeres generally activate damage response pathways that cause cell death or impairment of cellular function (2, 3). Evidence that cell replication is limited by shorter TL has been reported in vitro (37–40) through in vivo experimentation in animal models (41–43) and in cross-sectional and longitudinal observational studies (7–13, 44–48). On the other hand, whether inherited TL alters the risks of developing cancer is unsettled (20, 21, 49–53), so the implications of the association of paternal ages at reproduction with TL for cancer risk are unclear. Direct assessment of the health implications of paternal-age effects on descendants’ TL is lacking and awaits future study.

Together with this evidence, our findings lead us to the prediction that being born into a lineage in which recent male ancestors, particularly patrilineal ancestors, reproduced at later ages could have long-term health benefits. In particular, we expect changes in TL inherited at conception to have the largest effects on tissues and biological functions that are especially dependent upon rapid cellular proliferation, such as the immune system, the gastrointestinal tract, and skin (21). In contrast, cells that are not as dependent on proliferation, such as neurons or cardiomyocytes (21), might be less impacted by the age of reproduction in recent ancestors. Although some negative outcomes, such as miscarriages (54), are more common in offspring born to fathers of advanced age, the age-related lengthening of TL that we note here is present across the range of paternal and grandpaternal ages—not just at advanced ages.

The finding that telomere length may be extended cumulatively when multiple generations of male ancestors delay their ages of reproduction has potentially important evolutionary implications. Reproductive life span is a key evolutionary life-history parameter (e.g., 55–58), and this is particularly true in humans owing to the importance of late-life intergenerational transfers to reproduction and survival of offspring and grandoffspring (56, 59–61). In light of past work showing that blood TL predicts late-life survival (7–13), we believe that our findings point to a possible mechanism of plasticity in the pace of aging and senescent functional decline. Notably, as paternal ancestors delay reproduction, longer TL will be passed to offspring, which could allow life span to be extended as populations survive to reproduce at older ages. Having been born to an older father could signal that an individual is likely to grow up in a social and ecological context within which mortality rates are low and reproduction is likely to occur later in life, thus placing more of a premium on a durable long-lived body (21). However, the age at reproduction of one’s father alone is likely to be an imprecise signal of demographic conditions because factors like birth order will introduce additional variability. The high fidelity and cumulative multigenerational nature of the paternal age effect that we document are important in this regard because they suggest that TL could serve as a more robust signal of average paternal ages at reproduction over longer time frames (21). By integrating information about the average age at reproduction across multiple generations of ancestors, we speculate that the paternal age effect on TL could allow a unique form of transgenerational plasticity that modifies physiological function in response to a relatively stable cue of recent ancestral experience and behavior.

In sum, we find that grandpaternal age at the father’s birth is associated with longer TL in grandchildren, which is independent of, and additive to, the association of TL with the father’s age at the offspring’s birth. This evidence for cumulative, multigenerational lengthening of telomeres hints at a capacity for age-related changes in sperm TL to be transmitted with high fidelity across at least two generations. The implications of these findings for understanding heterogeneity in health and the pace of aging is an important question to be addressed by future research. We speculate that an effect of the age of paternal ancestors on TL could allow increases in life expectancy under demographic conditions of low mortality and delayed reproduction, when investment in a more durable and long-lived body is likely to reap higher fitness returns.

**Materials and Methods**

**Data Collection.** Data are from the Cebu Longitudinal Health and Nutrition Survey, a birth cohort study in Cebu City, the Philippines, that began with enrollment of 3,327 pregnant mothers in 1983–1984 (22). Mothers came from randomly selected rural and urban neighborhoods. Ages of the ancestor of the offspring were retrieved from household rosters (data available at http://www.cpc.unc.edu/projects/cebu), and as such there is missing age information reflecting ancestors who never lived with surveyed families. Venous blood samples were collected in 2005, when the offspring were 20.8–22.5 y old and the mothers were 35.7–69.3 y old. Written informed consent was obtained from all participants, and data and sample collection were conducted with approval and oversight from the Institutional Review Boards of...
Telomere Length Measurement. Automated and manual DNA extraction (Puregene, Gentra) was conducted on venous blood from 1,893 mothers and 1,779 offspring. DNA isolated in this fashion from blood is widely considered to be of leukocyte origin. Although this is no doubt mostly true, there are other possible cellular DNA sources, for example, extrachromosomal ribosomal RNA genes in Tetrahymena. J Mol Biol 120:33–53.


8. Cawthon RM, Smith KR, Olovnikov AM, Gentz K (1996) Telomeric DNA length measured by a monochrome multiplex quantitative PCR assay (23) with the following amplification conditions: 95 °C for 1.5 min at 95 °C; 20 s at 60 °C; 15 s at 98 °C; and 15 s at 74 °C with signal acquisition, followed by a melt curve for PCR product verification. Data were analyzed with a per-well efficiency calculation method (7, 67, 68) using LinRegPCR version 12.7 (69, 70). All telomere-to-single nucleotide (T/S) ratios were normalized to (divided by) the same control sample run with six replicates per 96-well plate. Intrarun correlations of the same samples (r = 0.92; n = 705; p < 0.0001) were comparable to a previous analysis (71). Samples with coefficients of variation (CVs) greater than 15% were rerun at least once, and 2.0% of all samples were dropped from the analysis due to CV < 15% not being achieved. Geometric mean CV of T/S ratios before exclusion was 5.7% and 5.6% after high-CV exclusion.

Statistical Analysis. Regression models were run with robust SEs because some of the regression models were heteroscedastic using the Breusch–Pagan–Cook–Weisberg test (“hettest” command in Stata). The household income variable, reflecting a probable multiplicative rather than additive effect, was logged. P values were combined using Fisher’s method (“metap” command in Stata). All tests were two-tailed with α = 0.05. All analyses were conducted using Stata 11.2.

ACKNOWLEDGMENTS. DNA extracts were provided by Karen Mohlke, Richard Cawthon, Justin Tackney, Katarina Nordfeldt, Klizia Salpea, Christine Ackermman, and Margrit Urbanek provided laboratory advice; and Jared Bradac, Yarrow Axford, and Zaneta Thayer gave feedback on earlier drafts of this manuscript. We thank the many researchers at the Office of Population Studies, University of San Carlos, for their central role in study design and data collection and the Filipino participants who provided their time and samples for this study. Laboratory analysis was funded by the National Science Foundation (Doctoral Dissertation Improvement Grant BCS-0962822) and the Ware–Gren Foundation (Grant 8111). Institutional support was received from Northwestern University. Data and sample collection were funded by National Institutes of Health Grants RR20649 and ES10126. D.T.A.E. was supported by a National Science Foundation Graduate Research Fellowship.

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